

## 應用機率模型作業 3 解答

Chapter2: #73, Chapter3: #24, #63

### 1.

(a) & (b).

$$\because N_i \sim \text{binomial}(n, p_i)$$

$$\therefore E(N_i) = np_i, \quad \text{Var}(N_i) = np_i(1 - p_i).$$

(c).

$$\because (N_i, N_j) \sim \text{multinomial}(n, p_i, p_j)$$

$$\therefore (N_i | N_j = n_j) \sim \text{binomial}\left(n - n_j, \frac{p_i}{1 - p_j}\right)$$

Hence,

$$\begin{aligned} \text{Cov}(N_i, N_j) &= \text{Cov}\left[E(N_i | N_j), N_j\right] = \text{Cov}\left[(n - N_j)\left(\frac{p_i}{1 - p_j}\right), N_j\right] = -\left(\frac{p_i}{1 - p_j}\right)\text{Var}(N_j) \\ &= -\left(\frac{p_i}{1 - p_j}\right)np_j(1 - p_j) = -np_ip_j. \end{aligned}$$

(d).

$$\text{Let } I_i = \begin{cases} 1, & \text{if the } i^{\text{th}} \text{ outcome do not occur.} \\ 0, & \text{otherwise.} \end{cases}$$

$$\therefore E\left(\sum_{i=1}^r I_i\right) = \sum_{i=1}^r E(I_i) = \sum_{i=1}^r P(N_i = 0) = \sum_{i=1}^r (1 - p_i)^n.$$

### 2.

(a).

Let  $N_{HT}$  denote the number of flips until at least one head and one tail have been flipped.

$N_1$  denote the number of flips needed “after the first flip”.

$$X = \begin{cases} 1, & \text{if the first outcome is head.} \\ 0, & \text{otherwise.} \end{cases}$$

Since  $(N_1 | X = 0) \sim \text{geometric}(p)$ , and  $(N_1 | X = 1) \sim \text{geometric}(1-p)$ .

$$\begin{aligned}\therefore E(N_{HT}) &= 1 + E(N_1) = 1 + EE(N_1 | X) = 1 + E(N_1 | X = 0) * (1-p) + E(N_1 | X = 1) * p \\ &= 1 + \frac{1-p}{p} + \frac{p}{1-p}.\end{aligned}$$

**(b) & (c).**

Let  $N_H$  denote the number of flips that land on heads.

$N_T$  denote the number of flips that land on tails.

$\therefore E(N_H | X = 0) = 1$ , and  $(N_H | X = 1) = (N_1 | X = 1) \sim \text{geometric}(1-p)$ .

$$\begin{aligned}\therefore E(N_H) &= EE(N_H | X) = E(N_H | X = 0) * (1-p) + E(N_H | X = 1) * p \\ &= 1-p + \frac{p}{1-p}.\end{aligned}$$

$$E(N_T) = E(N - N_H) = \left[ 1 + \frac{1-p}{p} + \frac{p}{1-p} \right] - \left[ 1-p + \frac{p}{1-p} \right] = \frac{1-p}{p} + p.$$

**(d).**

Let  $N_{2HT}$  denote the number of flips until at least two head and one tail have been flipped.

$$\begin{aligned}\therefore E(N_{2HT}) &= EE(N_{2HT} | X) = E(N_{2HT} | X = 0) * (1-p) + E(N_{2HT} | X = 1) * p \\ &= (1 + E(\text{until two head have been flipped})) * (1-p) + (1 + E(N_{HT})) * p \\ &= \left( 1 + \frac{2}{p} \right) * (1-p) + \left( 2 + \frac{1-p}{p} + \frac{p}{1-p} \right) * p.\end{aligned}$$

### 3.

**(a) & (b)**

Let  $I_i = \begin{cases} 1, & \text{if only one type } i \text{ in the final set.} \\ 0, & \text{otherwise.} \end{cases}$

$$\begin{aligned} E(I_i) &= P(I_i = 1) = EP(I_i = 1 | T) = \sum_{j=0}^{n-1} P(I_i = 1 | T = j) * P(T = j) = \frac{1}{n} \sum_{j=0}^{n-1} P(I_i = 1 | T = j) \\ &= \frac{1}{n} \sum_{j=0}^{n-1} \frac{1}{n-j}. \end{aligned}$$

The final equality follows because given that there are still  $n - j - 1$  uncollected types when the first type  $i$  is obtained, the probability starting at that point that it will be the last of the set of  $n - j$  types consisting of type  $i$  along with the  $n - j - 1$  yet uncollected types to be obtained is, by symmetry,  $1/(n - j)$ . Hence,

$$E\left(\sum_{i=1}^n I_i\right) = nE(I_i) = \sum_{j=1}^n \frac{1}{j}.$$